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The core-collapse rate from the Supernova Legacy Survey


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ABSTRACT

We use three years of data from the Supernova Legacy Survey (SNLS) to study the general properties of core-collapse and type Ia supernovae. This is the first such study using the “rolling search” technique which guarantees well-sampled SNLS light curves and good efficiency for supernovae brighter than $i' > 24$. Using host photometric redshifts, we measure the supernova absolute magnitude distribution down to luminosities 4.5 mag fainter than normal SNIa. Using spectroscopy and light-curve fitting to discriminate against SNIa, we find a sample of 117 core-collapse supernova candidates with redshifts $z < 0.4$ (median redshift of 0.29) and measure their rate to be larger than the type Ia supernova rate by a factor 4.5 ± 0.8(stat.) ± 0.6(sys.). This corresponds to a core-collapse rate at $z = 0.3$ of $(1.42 ± 0.3(stat.) ± 0.3(sys.)) × 10^{-4}$ yr$^{-1}$ (h$_{70}^{-3}$ Mpc$^{-3}$)

Key words. supernovae: general – stars: evolution – surveys

1. Introduction

The rate of supernova explosions is astrophysically important because it determines the rate at which heavy elements are dispersed into the interstellar medium, thereby constraining galactic chemical evolution. Since the progenitors of core-collapse supernovae (SNIa) are believed to be short-lived massive stars, the SNIa rate is expected to reflect the star-formation rate, in-
the majority of our SN candidates SNIa.

Since supernovae without spectroscopic identification, knowledge of the host redshift allows us to determine if the supernova four-band light-curves are consistent with the family of light curves typical of SNIa. The combination of spectroscopic and photometric typing will allow us to identify most SNIa. A relatively uncontaminated sample of SNcc is then defined as those supernovae not identified as SNIa. Use of the previously measured SNIa rate (Neill et al. 2006) will then allow us to derive the SNcc rate. The measurement will use only supernovae with redshifts \(<0.4\), beyond which the efficiency for detecting SNcc is too small to add significantly to the sample. This has the additional advantage that in this redshift range, the 615 nm Si II absorption feature is visible simplifying spectroscopic identification of SNIa.

The outline of this paper is as follows. Section 2 presents the light curve construction and event selection. Section 3 presents the characteristics of the supernova candidates. Section 4 defines the SNIa and SNcc candidate samples. Section 5 derives the relative SNIa and SNcc rates from which we deduce the SNcc rate. Section 6 concludes with a comparison of previous results.

Throughout, magnitudes are expressed in the AB system (Fukugita et al. 1996). A flat $\Lambda$CDM universe with $\Omega_M = 0.27$ is assumed.

2. Event selection

For this study, we performed a “deferred” search for transient events that was completely independent of the real-time search\(^1\) used to select supernovae for spectroscopy targets and for subsequent use in cosmological parameter analyzes. The details of the deferred search are given elsewhere (Bazin 2008; Bazin et al. 2009). We used SNLS observations of the four “deep” fields (D1, D2, D3, D4) from January 1st, 2003 to September 21, 2006. A reference image for each field and filter was constructed by co-adding the images from 20 good quality nights. The reference image was then subtracted from all science images of the same field and filter (after seeing-adjustment). In the $i'$ filter, the subtracted images from each lunation were combined to form one “stacked” image per lunation and stellar objects were searched for on each of these stacks. Approximately 300 000 objects were found, mostly spurious detections due to saturated signals from bright objects. Four-filter light curves for these objects were then

\(^1\) http://legacy.astro.utoronto.ca/
obtained from individual subtracted images by differential photometry with PSF fitting, imposing the position found on the $i'$ stack. Fluxes were calibrated using the set of SNLS tertiary standards (Astier et al. 2006).

The event selection criteria applied on the detected light curves are described in detail in Bazin et al. (2009). Spurious detections were mostly eliminated by requiring that the light curves in $i'$ and $r'$ have at least three successive photometric points with fluxes above 1σ from base line and their dates of maximum flux should be within 50 days from each other. Light curves corresponding to detections near stars as identified in our reference images were also discarded. Accepted light curves were fit with the phenomenological form

$$f(t) = A \frac{e^{-\frac{t-t_{0}}{\tau_{\text{rise}}}}}{1 + e^{\frac{t-t_{0}}{\tau_{\text{fall}}}}} + B$$

while this form has no particular physical motivation, it is sufficiently general to fit the shape of all types of supernovae. Long-term variable objects (such as AGNs) were rejected by comparing the $\chi^2$ of the light curve fit using (1) with fits to a constant flux, and only accepting objects for which the phenomenological model is a substantially better fit. In addition, we require that the light-curve be consistent with a time-independent flux before and after the main variation as fitted by (1). The precise cuts were defined with the help of synthetic SNIa light curves and selected real light curves which have been confirmed by spectroscopy as type Ia or core-collapse SNe. Finally, good time sampling criteria were applied, i.e. requiring at least one pre-max epoch within 30 days and one post-max epoch within 60 days of the date of maximum flux in the $i'$ and $r'$ filters, and at least two epochs in that time interval in the $g'$ and $z'$ filters. A set of 1462 events was thus retained.

Light-curves for three events are shown in Figs. 1–3. The first shows a typical spectroscopically-confirmed SNIa with spectroscopic redshift $z = 0.332$ and the second a typical spectroscopically-confirmed SNcc with spectroscopic redshift $z = 0.328$. The third is one of the faintest events to be used in Sect. 5 to measure the core-collapse rate. Its peak magnitude is $i' = 24.1$, as fitted by (1).

To identify host galaxies for the events, we used the photometric galaxy catalog of Ilbert et al. (2006). The host for an event was chosen to be the galaxy with the smallest distance, $r$, between the event and the galaxy center in units of the galaxy’s effective radius, $r_{\text{gal}}$, defined as the half-width of the galaxy in the direction of the event. The value of $r_{\text{gal}}$ was defined by the $A$, $B$ and $\theta$ SExtractor parameters (Bertin & Arnouts 1996). The match was considered successful if the host was at a distance $r < 5r_{\text{gal}}$. This choice was a compromise between host finding efficiency and accidental mismatching. Of the 1462 selected events, 1329 (91%) have matched hosts and of these 1207 (91%) were targeted during the period considered here including events which appear to be associated with low redshift galaxies. After elimination of these events, we were left with 221 events.

The efficiency of the event selection procedure was calculated by treating simulated supernovae with the same procedure. Supernovae added to real $i'$-band images were used to test the initial detection stage in $i'$. The efficiency of the subsequent event selection cuts was calculated by applying them to light curves generated by a Monte-Carlo simulation that takes into account the photometric resolution and the observing sequence. The resulting efficiency is a function of the maximum fluxes in the four bands and the associated time scales. However, to good approximation the efficiency is simply a function of the maximum in the $i'$ band. The efficiency is shown in Fig. 5 for SNIa and for long SNcc ($\tau_{\text{fall}} = 100$ days). In both cases, the efficiency is relatively $i'$-independent at a value of $\sim 0.8$ for $i' < 23$ at which point it starts to decline, reaching 0.4 at $i' = 24.3$.

The performance of our selection pipeline was checked by comparing it with the results of the SNLS real-time pipeline used to select spectroscopy targets. A total of 340 supernovae were targeted during the period considered here including events
as faint as $i' = 24.4$. Of these, all but two were found on the $i'$ stacked images. (The two lost events were outside the reference images.) Of the 338 events, 295 passed our selection criteria. The loss of the 43 events was due to our time sampling criteria which is more restrictive than the real-time criteria.

3. Event characteristics

Figure 6 shows the $i'$ Hubble diagram for the 221 events with host photometric redshifts <0.4. Events that are spectroscopically identified as SNIa or SNcc (SNIi, SNIfb, SNIc) are marked. Also marked are photometrically identified SNccs. As expected for a sample dominated by SNIa and SNcc, the faint events have relatively fewer early-type hosts (19/152) compared to 24/69 for the bright events. Figure 7 shows $\Delta M_{570}$ as a function of redshift. Identified SNIa are centered near $\Delta M_{570} < 0.75$ containing most spectroscopically confirmed SNIa. The characteristics of the events as a function of $\Delta M_{570}$, shown in Figs. 8–10, are broadly consistent with those expected for SNIa and SNcc. Figure 8 shows $\Delta M_{570}$ as a function of the host type (Ilbert et al. 2006). As with low redshift SNIa (Richardson et al. 2002), about half (47/108) the faint events have $\tau_{\text{fall}}/(1+z) > 50$ days, characteristic of plateau SNIi and significantly longer than fall times for SNIa, 20 < $\tau_{\text{fall}}/(1+z) < 30$ days. Finally, Fig. 10 shows the color–magnitude diagram using the AB magnitude at 450 nm in the rest frame:

$$m_{450} \equiv (4z - 0.4)r' + (1.4 - 4z)g'.$$

$$\Delta M_{570} \equiv m_{570} - 2.5 \log \left(1+z \right) d(z)^2 - C$$

where

$$d = \int^{\infty}_{0} \frac{dz'}{\sqrt{\Omega_m(1+z)^3 + (1-\Omega_m)}} = \frac{d_{L}}{(c/H_0)(1+z)}.$$
The reasonable performance of the photometric classification is demonstrated by the fact that only seven of the 46 spectroscopic SNIa were not selected and only one of the 24 spectroscopic SNcc was selected. The photometric classification also selected 14 events that had not been classified spectroscopically as SNIa. As detailed below, the lack of spectroscopic confirmation was generally due to an insufficient supernova signal over the galactic background or to lack of telescope time to obtain a spectrum.

As our nominal SNIa sample, we choose the 46 events that were spectroscopically identified as SNIa or “not SNIa” for events classified spectroscopically as SNIa, SNcc and “ambiguous”, as well as for events for which no spectrum was obtained. The table contains only those events that will be used for rate measurements in the next section, i.e. those with $0.05 < z < 0.4$ and $m_{570} < 24.1$ (Eq. (2)).

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character: three had photometric redshifts significantly different from the spectroscopic redshifts causing the SALT fit to be very poor; one event had an extreme color parameter falling outside our cuts; three events had a small number of poor photometric points causing the fits to fail our χ² cut.

We now consider contamination of the 14 SNIa that have only photometric confirmation. Of these, four events have spectroscopy that was of insufficient signal-to-noise to determine the SN type. The remaining 10 events had no spectra either because the event was discovered too late or because the estimated signal-to-noise was insufficient (local flux increase <20%). Only one of the 14 events was judged “unlikely” to be a SNIa by the spectroscopy target selection group, but the full light curve indicates that it is consistent with being a normal SNIa. We therefore have no evidence that the 14 events are contaminated with SNcc. However, we have no good template catalog of bright SNIa to evaluate the probability that a bright SNIa passes our SNIa photometric selection. We therefore conservatively assign a systematic one standard deviation upper limit of 7 events to contamination of the SNIa sample with SNcc.

While we have no evidence that the SNIa sample is contaminated with SNcc, it is certain that the SNcc sample is contaminated by sub-luminous SNIa. We will evaluate this contamination in the next section.

5. The core-collapse rate

In this section, we will derive the SNcc rate using events with 0.05 < z < 0.4 and m_{570} < 24.1. The cut on m_{570} is used to ensure that only events with good detection efficiency are used. The requirement that z > 0.05 eliminates one event at z = 0.04. The uncertainty in ΔM_{570} is σ(ΔM_{570}) ~ 2σ_r/z so the event with z = 0.04 has σ(ΔM_{570}) > 1 and we prefer to eliminate it. (In fact, this event is a spectroscopic outlier with z_{spec} = 0.247.) With the m_{570} and redshift cuts, we are left with 177 events, 60 of which are spectroscopically or photometrically identified SNIa.

From these numbers, we will derive the SNcc rate as follows. We first assign weights to the observed events that take into account detection efficiency and the volume over which the event could be detected by SNLS. Because of their intrinsic faintness, this will significantly increase the number of SNcc candidates to 287. We next evaluate two effects that can change the number of SNcc candidates relative to SNIa candidates. The first is simple spectral or photometric misidentification. The second comes from the use of host photometric redshifts which, we will see, has a slight tendency to increase the number of SNIa candidates relative to SNcc candidates. Using the corrected number of candidates, we can calculate the SNcc rate relative to the SNIa rate. By adopting the previously measured SNIa rate, we then derive the SNcc rate for luminosities within 4.5 mag of normal SNIa. Finally, we estimate the total SNcc rate taking into account the decrease in the number of observed supernovae due to extinction by dust in the host galaxy.

5.1. Event weights

The observed distribution of ΔM_{570} is the histogram shown in Fig. 11 for the 177 events with z < 0.4 and m_{570} < 24.1. In order to derive the true distribution of ΔM_{570} for events with z < 0.4, this distribution must be corrected for the i'-dependent detection efficiency and, more importantly, for the fact that an event with absolute magnitude ΔM_{570} can be seen only up to a redshift, z_{max} defined by

\[ \sqrt{1 + z_{max}} = 10^{0.2(m_{570, max} - c - ΔM_{570})} \]  

where d(z) is defined by (4) and m_{570, max} = 24.1 is the maximum accepted magnitude. Events with z_{max} < 0.4 (i.e. ΔM_{570} > 1.7) must be given a weight, W > 1. The weight takes into account, most importantly, the fact that SNLS detects them over a smaller volume. Of secondary but non negligible importance is the fact that the SNcc rate is believed to increase with redshift, R_{cc} ∝ (1 + z)^{α} with α ∼ 3.6 to reflect the increasing star-formation rate with redshift. We are therefore sensitive to intrinsically faint supernovae only in a redshift range where rate is small. We correct for this with α = 3.6 appropriate for SNcc because our SNIa candidates are all bright enough to have z_{max} > 0.4. Finally, because of cosmological time dilation, the SNLS observing time is proportional to (1 + z)^{-1}. The total weight, W(ΔM_{570}, i'), is therefore given by

\[ W^{-1} ∝ \frac{\epsilon(i')}{\epsilon(i' = 21)} \int_0^{z_{max}} \frac{d\epsilon(d(z)^3(1 + z)^{y-1}}{\sqrt{\Omega_m(1 + z)^3} + (1 - \Omega_m)} \]  

where \(\epsilon(i')\) is the event detection efficiency. The factor of proportionality is chosen so that W(z_{max} = 0.4, i' = 21) = 1. Without the factors of (1 + z) in (4) and (7), the weight would be simply the product of \(\epsilon(i' = 21)\)/\(\epsilon(i')\) and the euclidean volume ratio (0.4/z_{max})^3. The factors of (1 + z) correct for the redshift evolution of the volume element, the exposure time, and the SNcc rate.

Weighting individual events gives the corrected ΔM_{570} distribution shown by the data points and error bars in Fig. 11. All of the 60 SNIa candidates have weights near unity. Because of their faintness, many of the 117 SNcc candidates have W > 1 and the corrected number of SNcc candidates is 287 ± 40 (statistical error only).

5.2. Corrections for misidentification and redshift migration

In this section we correct the raw number of SNIa and SNcc candidates for two effects that can affect their numbers: type misidentification (summarized in Table 2) and redshift migration due to the use of photometric redshifts.
The first shift in the SNIa-SNcc ratio is due to SNIa that are incorrectly identified as SNcc. We divided this correction into that for “sub-luminous” SNIa and normal SNIa. Sub-luminous SNIa (Li et al. 2001) have a mean magnitude 1.5 mag below the mean magnitude for normal SNIa and account for 16 ± 6% of SNIa. None are found in our sample since both selection for spectroscopy and photometric selection aimed at finding normal SNIa. We therefore add (subtract) 9 ± 5 events to the SNIa (from the SNcc) samples. For normal SNIa, we must correct for events that were neither spectroscopically nor photometrically selected. From Table 1, of the 46 spectroscopically confirmed events, only 7 were not photometrically selected. This gives an inefficiency of 7/46 = 0.15 for photometric identification of spectroscopically confirmed SNIa. To the 14 SNIa candidates relying solely on photometric selection, we can therefore add 0.15 × 14 = 2 events and subtract the same number from the SNcc.

As discussed in Sect. 4, we make no correction for SNIa incorrectly identified as SNIa but assign a systematic one standard deviation upper limit of 7 events to contamination of the SNIa sample with SNcc.

Contamination with non-supernova events is expected to be unimportant. The scan of events resulted in the elimination of only six AGN-like events and the identification of four additional events that were judged uncertain. We adopt four events as our one standard deviation upper limit on AGN contamination of the SNcc sample.

Finally, we correct for redshift migration (Eddington bias), an effect that comes from our use of photometric redshifts with a modest resolution of σz = 0.04. Because there are more supernovae at high redshift than at low redshift, the main effect of this resolution is for high redshift supernovae to migrate below the z = 0.4 cutoff. If there were no cut m570 < 24.1, this would increase the number of SNIa and SNcc candidates by the same factor. The fact that SNcc are fainter than SNIa means that migrating SNcc are less likely to satisfy m570 < 24.1 than migrating SNIa. We have used a Monte Carlo simulation to estimate this effect. The simulation generates events with a realistic redshift and M570 distribution and uses the observed spectroscopic-photometric redshift pairs from Fig. 4 to assign photometric redshifts. Outliers in this plot are used so the simulation takes into account catastrophic redshifts. Counting weighted simulated events indicates that the migration makes the measured SNcc-SNIa rate ratio (15 ± 4)% less than the real rate ratio. The statistical error comes from the limited number of redshift pairs we have used for the simulation. The measured SNcc-SNIa rate will therefore be multiplied by a factor 1.15 to take into account this effect.

5.3. The SNcc-SNIa relative rate

The corrections for the number of events shown in Table 2 give an increase of 11.5 ± 5 SNIa candidates, and a corresponding decrease in the number of SNcc candidates. The corrected relative rate for z < 0.4 is therefore

\[
\frac{R_{cc}(\Delta M_{570} < 4.5)}{R_{cc}} = \frac{287 - 11}{60 + 11} \times 1.15
\]

where the factor 1.15 takes into account redshift migration. The ratio is for z < 0.4 corresponding to an expected mean redshift of 0.306 for a rate proportional to (1 + z)² and a mean of 0.313 for a rate proportional to (1 + z)⁰. Our sample of 60 SNIa has a mean redshift of 0.30 ± 0.01, consistent with expectations for a complete (volume limited) sample.

The systematic error in \(R_{cc}/R_{cc}\) includes those due to the corrections from the previous section as well as three additional systematic uncertainties which we add in quadrature.

The first additional systematic concerns the uncertainty in the relative efficiencies for SNIa and the fainter SNcc. To avoid large uncertainties, we have used only events with m570 < 24.1 where the efficiency is high. With this cut, there is only a 10% difference in the SNcc rate calculated with the nominal efficiencies and that calculated assuming a magnitude-independent efficiency. We adopt 10% as the nominal systematic error from this source. To check that there is no significant uncorrected event loss near the magnitude cut, we verified that the derived SNcc rate does not depend significantly on the position of the magnitude cut. For example, using m570 max = 23.6, the number of events SNcc candidates is reduced from 117 to 82. After weighting, this is increased to 334±80 consistent with the 287±45 event found using m570 max = 24.1. (Most of the increase comes from the two events with \(\Delta M_{570} > 3.5\) which are given greater weights with m570 max = 23.6.)

The second systematic concerns the star-formation rate. The corrected differential rate was calculated assuming that the SNcc rate is proportional to \((1 + z)^α\) with α = 3.6 according to Hopkins & Beacom (2006). These authors do not cite an uncertainty for α but inspection of the data indicates that α = 3.6 ± 1.0 is reasonable. This corresponds to a 10% systematic uncertainty in the SNcc rate.

The final systematic concerns our requirement that a host galaxy be found and that a redshift be given in the Ilbert et al. (2006) catalog. This requirement could conceivably favor SNcc or SNIa. For spectroscopically identified supernovae with z < 0.4, a host galaxy is generally found but a redshift may not be given in the catalog. For 41 spectroscopic SNIa with spectroscopic redshifts <0.4 that were found in the deferred search, only 4 have no host redshift while for the 36 spectroscopic SNcc there are only 3 with no host redshift. Thus, we see no difference in host-redshift measurement efficiency at the 5% level and we adopt this as the systematic uncertainty.

5.4. The SNcc rate

To derive a value of Rcc we adopt the value of Rcc measured byNeill et al. (2006) at \(z = 0.5: 0.42 \times 10^{-4} \text{yr}^{-1}(h_{70}^{-1} \text{Mpc})^{-3}\). Our measurement of Rcc/Rcc is effectively at \(z = 0.3\) and we adopt a SNcc rate at this redshift of 0.315×10⁻⁴ yr⁻¹(h₇₀⁻¹ Mpc)⁻³ calculated assuming Rcc = \((1 + z)^2\). This gives a SNcc rate within 4.5 mag of normal SNIa of

\[
\frac{R_{cc}(\Delta M_{570} < 4.5)}{10^{-4} \text{yr}^{-1}(h_{70}^{-1} \text{Mpc})^{-3}} = 1.42 \pm 0.30 \text{ (stat.)} \pm 0.32 \text{ (sys.)}
\]

we have added the statistical and systematic uncertainties of \(R_{cc}/R_{cc}\) and of \(R_{cc}\) separately in quadrature though not including
the systematic uncertainty in \( R_{\text{fc}} \) due to sub-luminous and absorbed supernovae because it is already included in the uncertainty in \( R_{\text{fc}} / R_{\text{fc}} \).

With the determination of the total SNcc rate, we can give an absolute differential rate per unit magnitude for SNcc. It is shown as the right-hand scale in Fig. 12. The rate is measured down to luminosities \( 4.5 \) mag fainter than normal SNIa. It should however be emphasized that there are only two events with \( \Delta M_{570} \geq 3.5 \). One of them has spectroscopic confirmation and the spectroscopic redshift, \( z = 0.131 \), is in good agreement with the host photometric redshift, \( z = 0.119 \). The other event has a host spectroscopic redshift \(^2 z = 0.0815 \) in good agreement with the host photometric redshift used here, \( z = 0.095 \). Thus, we have no indication that these two events are higher luminosity events that have migrated from high redshift.

To estimate a total rate for SNcc we need to estimate the number of SNcc with \( \Delta M_{570} \geq 4.5 \) either because they are intrinsically faint (e.g. SN1987A, \( \Delta M_{570} \sim 5.5 \)) or because of high host extinction. SNLS obviously cannot say anything about intrinsically faint supernovae. However, by adopting a host galaxy extinction model, we can estimate the number of SNcc that have intrinsic luminosities within our range of sensitivity but that are lost because of high host extinction. We have used the results of Hatano et al. (1998) who give (their Table 1 and Fig. 1) the distribution of \( A_B \) as a function of host inclination angle. This can be converted to a distribution of absorption at 570 nm and convoluted with the pre-absorption distribution of \( M_{570} \). For example, if we model the intrinsic SNcc magnitude distribution shown as the dashed line in Fig. 12, then the SNcc host extinction model of Hatano et al. (1998) predicts the distribution shown by the solid line in the figure. With this model, 15% of SNcc have \( \Delta M_{570} > 4.5 \). Our estimated total rate is then increased to \( R_{\text{fc}} = 1.63 \times 10^{-4} \text{yr}^{-1} (h_{70}^{-1} \text{Mpc})^{-3} \). In our model, most of the events with \( \Delta M_{570} > 4.5 \) are highly absorbed so our estimate should be considered a lower limit on the SNcc rate that ignores supernovae that are intrinsically fainter than \( \Delta M_{570} = 4.5 \).

6. Discussion

Figure 13 summarizes the published measurement of the SNcc rate. All data is consistent with a rate that increases with redshift like the SFR \( \propto (1 + z)^{3.6} \). It should be emphasized that the previous measurements use quite different detection and analysis procedures. We therefore refrain from drawing any quantitative conclusions about the redshift dependence of the SNcc rate.

Our results will be improved in the future with the addition of two more years of SNLS data, and with the use of host spectroscopic redshifts that we are in the process of obtaining.

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\(^2 \text{http://nedwww.ipac.caltech.edu/index.html} \)